

A Regularizing Ensemble Kalman Method for PDE-constrained Inverse Problems

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1 Introduction

2 Numerical Investigation of the Scheme

3 Applications

1 Introduction

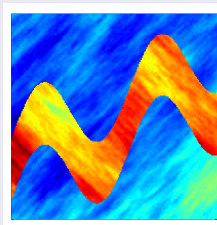
2 Numerical Investigation of the Scheme

3 Applications

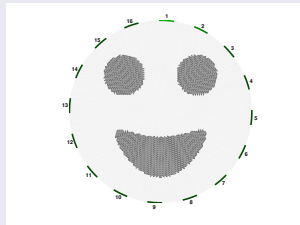
Inferring/estimating functions which are inputs for a PDE model, given measurements/observations form the output.

PDE-constrained applications

Porous media flow



Electrical Impedance Tomography



M. A. Iglesias

A regularizing ensemble Kalman method for PDE-constrained inverse problems.
to appear in Inverse Problems, 2015. <http://arxiv.org/abs/1505.03876>



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Iterative regularization for ensemble data assimilation in reservoir modeling
Computational Geosciences, (2015) 19:177-212

Let $\mathcal{G} : X \rightarrow \mathbb{R}^J$.

Forward Model

Given $u \in X$ compute

$$y = \mathcal{G}(u).$$

Let $\eta \in \mathbb{R}^J$ be a realization of an observational **noise**.

Inverse Problem

Given $y \in \mathbb{R}^J$ find $u \in X$:

$$y = \mathcal{G}(u) + \eta.$$

Forward groundwater flow model:

Forward Problem: Darcy flow

$$\begin{aligned} -\nabla \cdot \kappa \nabla p &= f && \text{in } D, \\ -\kappa \nabla p \cdot n &= B_N && \text{in } \Gamma_N, \\ p &= B_D && \text{in } \Gamma_D \end{aligned}$$

where $\partial D = \Gamma_N \cup \Gamma_D$.

$$u = \log(\kappa(x)) \in X \equiv L^\infty(D) \longrightarrow \mathcal{G}(u) = \{p(x_i)\}_{i=1}^J \in \mathbb{R}^J$$

Inverse Problem

Given $y \in \mathbb{R}^J$ find $u \in X$:

$$y = \mathcal{G}(u) + \eta.$$

Prior

Probabilistic information about u *before* data is collected:

$$\mu_0(u) = \mathbb{P}(u)$$

Likelihood

Since $y = \mathcal{G}(u) + \eta$, if $\eta \sim N(0, \Gamma)$, then $\mathbb{P}(y|u) = N(\mathcal{G}(u), \Gamma)$. Then (Γ -weighted) **model-data misfit** Φ is the negative log-likelihood:

$$\Phi(u; y) = \frac{1}{2} \left\| \Gamma^{-1/2} (y - \mathcal{G}(u)) \right\|^2$$

Posterior

Probabilistic information about u *after* data is collected:

$$\mu^y(u) = \mathbb{P}(u|y).$$

$$\frac{\mu^y(u)}{\mu_0(u)} \propto \exp \left(- \Phi(u; y) \right)$$

Posterior

Probabilistic information about u *after* data is collected:

$$\mu^y(u) = \mathbb{P}(u|y).$$

$$\frac{\mu^y(u)}{\mu_0(u)} \propto \exp\left(-\Phi(u; y)\right)$$

Challenge

To explore the probability measure μ^y .

- \mathcal{G} is highly nonlinear; μ^y cannot be characterized with a few parameters.
- The problem is high dimensional (X is discretized with 10^6 - 10^9 cells).
- Standard sampling methods for Bayesian inference do not work.
- Infinite-dimensional Bayesian framework [Stuart, 2010]; MCMC method for functions (pcn-MCMC) [Cotter, et-al, 2013].
- Well-known for continuous \mathcal{G} .

The Classical (deterministic) formulation of the Inverse Problem

Given data $y \in Y$ find

$$u = \arg \min_{u \in X} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2 \rightarrow \min$$

For most PDE-constrained applications $\mathcal{G} : X \rightarrow \mathbb{R}^M$ is compact (unless X is finite dimensional)

Lack of continuity (lack of stability) with respect to the data

We can construct a sequence $u_n \in X$ such that

$$u_n \not\rightarrow u \quad \text{but} \quad \mathcal{G}(u_n) \rightarrow \mathcal{G}(u)$$

If we want to compute the minimizer above with standard optimization we may observe semiconvergence behavior [Kirsch, 1996]

Regularization Approaches (for nonlinear operators)

- Regularize-then-compute (e.g. Tikhonov, TSVD)
- Compute while regularizing (**Iterative Regularization**) [Kaltenbacher, 2010]
 - *regularizing Levenberg-Marquardt*
 - *Landweber iteration*
 - *truncated Newton-CG*
 - *iterative regularized Gauss-Newton method*

Introduce noise level

$$\|\Gamma^{-1/2}(y - \mathcal{G}(u^\dagger))\| \leq \eta$$

Regularization

Construct an approximation u^η that is stable, i.e. such that

$$u^\eta \rightarrow u \quad \text{as} \quad \eta \rightarrow 0$$

where

$$\mathcal{G}(u) = \mathcal{G}(u^\dagger)$$

Consider $\mu_0(u) = \mathbb{P}(u) = \mathcal{N}(0, C)$ the prior on u and

$$y = \mathcal{G}(u) + \xi, \quad \xi \sim \mathcal{N}(0, \Gamma)$$

The Bayesian Inverse Problem

Characterize the posterior $\mu^y(u) = \mathbb{P}(u|y)$:

$$\frac{\mu^y(u)}{\mu_0(u)} \propto \exp\left(-\frac{1}{2}\|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2\right) \quad y^{(j)} = y + \eta^{(j)} \sim \mathcal{N}(0, \Gamma)$$

Ensemble Approximating the Bayesian posterior $\mu^y(u) = \mathbb{P}(u|y)$

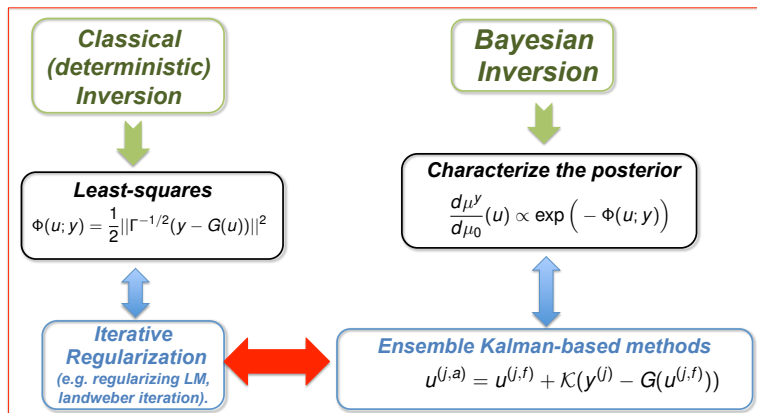
Randomizing least-squares, e.g.

$$\frac{1}{2}\|\Gamma^{-1/2}(y^{(j)} - \mathcal{G}(u))\|^2 \rightarrow \min \quad y^{(j)} = y + \eta^{(j)} \sim \mathcal{N}(0, \Gamma)$$

or, for example, Randomized Maximum Likelihood

$$\frac{1}{2}\|\Gamma^{-1/2}(y^{(j)} - \mathcal{G}(u^{(j)}))\|^2 + \|C^{-1/2}(u - u^{(j)})\|_X^2 \rightarrow \min \quad u^{(j)} \sim \mathcal{N}(0, C)$$

Overview of this work



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Bayesian Inverse Problem

Given a prior $\mu_0(u)$ of on u and data

$$y = \mathcal{G}(u) + \eta \quad \eta \sim N(0, \Gamma)$$

find $\mu(u) = \mathbb{P}(u|y)$.

$$\mu(u) \propto \mu_0(u) \exp \left\{ -\frac{1}{2} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2 \right\}$$

Define

$$z = \begin{pmatrix} u \\ \mathcal{G}(u) \end{pmatrix}, \quad y = Hz + \eta, \quad H = (0, I)$$

Alternative Bayesian Inverse Problem

Given a prior on z , $\mu_0(z)$ and data y , find $\mu(z) = \mathbb{P}(z|y)$

$$\frac{\mu(z)}{\mu_0(z)} \propto \exp \left\{ -\frac{1}{2} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2 \right\}$$

Bayesian Inverse Problem

Given a prior on z , $\mu_0(z)$ and data y , find $\mu(z) = \mathbb{P}(z|y)$

$$\mu(u) \propto \mu_0(u) \exp \left\{ -\frac{1}{2} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2 \right\}$$

Construct an initial ensemble

$$z_0^{(j,f)} = \begin{pmatrix} u_0^{(j)} \\ \mathcal{G}(u_0^{(j)}) \end{pmatrix}, \quad \{u_0^{(j)}\}_{j=1}^{N_e} \sim \mu_0$$

Compute mean and covariance

$$\bar{z}^f = \frac{1}{N_e} \sum_{j=1}^{N_e} z^{(j,f)} \quad C^f = \frac{1}{(N_e - 1)} \sum_{j=1}^{N_e} z^{(j,f)} (z^{(j,f)})^T - \bar{z}^f (\bar{z}^f)^T$$

Gaussian Approximation: $\mu_0(z) = N(\bar{z}^f, C^f)$,

Since $\mu_0(z) = N(\bar{z}^f, C^f)$, then

$$\mu(u) \propto \mu_0(u) \exp \left\{ -\frac{1}{2} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2 \right\} = N(\bar{z}^a, C^a)$$

where

$$\bar{z}^{(a)} = \bar{z}^{(f)} + \underbrace{C^f H^T (H C^f H^T + \Gamma)^{-1}}_K (y - H \bar{z}^{(f)})$$
$$C^a = (I - K) C^f$$

Updated each ensemble according to

$$z^{(j,a)} = z^{(j,f)} + C^f H^T (H C^f H^T + \Gamma)^{-1} (y^{(j)} - H z^{(j,f)})$$

with

$$y^{(j)} = y + \eta^{(j)}, \quad \eta^{(j)} \sim N(0, \Gamma)$$

Ensemble Kalman Smoother

Updated each ensemble according to

$$z^{(j,a)} = z^{(j,f)} + C^f H^T \left(H C^f H^T + \Gamma \right)^{-1} (y^{(j)} - H z^{(j,f)})$$

Claim 1: $\{z^{(j,a)}\}_{j=1}^{N_e} \approx \mathbb{P}(z|y)$.

Recall that $z = \begin{pmatrix} u \\ \mathcal{G}(u) \end{pmatrix}$. Then,

$$u^{(j)} = u_0^{(j)} + C^{uw} (C^{ww} + \Gamma)^{-1} (y^{(j)} - G(u_0^{(j)}))$$

$$C^{uw} = \frac{1}{(N_e - 1)} \sum_{j=1}^{N_e} (u_0^{(j)} - \bar{u}_0) (\mathcal{G}(u_0^{(j)}) - \mathcal{G}(\bar{u}_0))^T$$

$$C^{ww} = \frac{1}{(N_e - 1)} \sum_{j=1}^{N_e} (\mathcal{G}(u_0^{(j)}) - \mathcal{G}(\bar{u}_0)) (\mathcal{G}(u_0^{(j)}) - \mathcal{G}(\bar{u}_0))^T$$

Claim 2: $\{u^{(j)}\}_{j=1}^{N_e} \approx \mathbb{P}(u|y)$.

Observation: $\{u^{(j)}\}_{j=1}^{N_e} = \mathbb{P}(u|y)$ if \mathcal{G} is linear and μ_0 is Gaussian.

It kind of works....sometimes.

$$u^{(j)} = u_0^{(j)} + C^{uw}(C^{ww} + \Gamma)^{-1}(y^{(j)} - \mathcal{G}(u_0^{(j)}))$$

Underestimates the uncertainty

Iterative smoothers

$$u_n^{(j)} = u_{n-1}^{(j)} + C_{n-1}^{uw}(C_{n-1}^{ww} + \Gamma)^{-1}(y^{(j)} - \mathcal{G}(u_{n-1}^{(j)}))$$

Overestimates the uncertainty

Ad-hoc fixes of Iterative smoothers

$$u_n^{(j)} = u_{n-1}^{(j)} + \rho \circ C_{n-1}^{uw}(C_{n-1}^{ww} + \alpha\Gamma)^{-1}(y^{(j)} - \mathcal{G}(u_{n-1}^{(j)}))$$

ρ is a localization matrix and α is an inflation parameter

Understanding the iterative ensemble smoother with iterative regularization

Suppose we are interested in solving

$$u = \arg \min_{u \in X} \|\Gamma^{-1/2}(y - \mathcal{G}(u))\|^2$$

where X has norm $\|C^{-1/2} \cdot\|$

Levenberg-Marquardt

u_{n+1} iteration level is given by

$$u_{n+1} = u_n + \arg \min_{v \in X} \|\Gamma^{-1/2}(y - \mathcal{G}(u_n) - D\mathcal{G}(u_n)v)\|^2 + \alpha \|C^{-1/2}v\|_X^2$$

After some computations

$$u_{n+1} = u_n + C D\mathcal{G}^*(u_n)(D\mathcal{G}(u_n) C D\mathcal{G}^*(u_n) + \alpha \Gamma)^{-1}(y - \mathcal{G}(u_n))$$

Hanke proposed a way to select α and a stopping criteria ($\delta =$ noise level):

$$\|\Gamma^{-1/2}(y - \mathcal{G}(u_n))\| \approx \delta$$

so that

$$u_{n+1} = u_n + C D\mathcal{G}^*(u_n)(D\mathcal{G}(u_n) C D\mathcal{G}^*(u_n) + \alpha\Gamma)^{-1}(y - \mathcal{G}(u_n))$$

generates a stable approximation to the solution of the classical inverse problem.

Theorem [Hanke 1997]

The LM scheme terminates after a finite number of iterations n^* and

$$u_{n^*} \rightarrow u \quad \text{as} \quad \eta \rightarrow 0 \quad (\text{where } \mathcal{G}(u) = \mathcal{G}(u^\dagger))$$

u^\dagger is the truth

Regularizing ensemble Kalman method

Consider an initial ensemble $\{u_0^{(j)}\}_{j=1}^{N_e} \subseteq X$

Linearize around the ensemble mean $\bar{u}_n \equiv \frac{1}{N_e} \sum_{j=1}^{N_e} u_n^{(j)}$

$$w_n^{(j)} \equiv \mathcal{G}(u_n^{(j)}) \approx \mathcal{G}(\bar{u}_n) + D\mathcal{G}(\bar{u}_n)(u_n^{(j)} - \bar{u}_n)$$

$$C_n^{uu} D\mathcal{G}^*(\bar{u}_n)v \approx C_n^{uw} v \qquad D\mathcal{G}(\bar{u}_n)C_n^{uu} D\mathcal{G}(\bar{u}_n)^* v \approx C_n^{ww} v$$

Recall the update formula for the regularizing LM scheme

$$u_{n+1} = u_n + C D\mathcal{G}^*(u_n)(D\mathcal{G}(u_n) C D\mathcal{G}^*(u_n) + \alpha\Gamma)^{-1}(y - \mathcal{G}(u_n))$$

Replace

$$\begin{aligned} u_n &\implies \bar{u}_n, \\ C D\mathcal{G}^*(u_n) &\implies C_n^{uu} D\mathcal{G}^*(\bar{u}_n) \approx C_n^{uw} \\ D\mathcal{G}(u_n) C D\mathcal{G}^*(u_n) &\implies D\mathcal{G}(\bar{u}_n)C_n^{uu} D\mathcal{G}^*(\bar{u}_n) \approx C_n^{ww} \end{aligned}$$

Update formula for the mean of the ensemble

$$\bar{u}_{n+1} = \bar{u}_n + C_n^{uw} (C_n^{ww} + \alpha\Gamma)^{-1} (y - \bar{w}_n)$$

where $\bar{w}_n \equiv \frac{1}{N_e} \sum_{j=1}^{N_e} \mathcal{G}(u_n^{(j)})$.

We propose to update each ensemble in a consistent fashion

$$u_{n+1}^{(j)} = u_n^{(j)} + C_n^{uw} (C_n^{ww} + \alpha\Gamma)^{-1} (y^{(j)} - \mathcal{G}(u_n^{(j)}))$$

Selection of α :

$$\rho \|\Gamma^{-1/2} (y - \bar{w}_n)\|_Y \leq \alpha \|\Gamma^{1/2} (C_n^{ww} + \alpha\Gamma)^{-1} (y - \bar{w}_n)\|_Y$$

Stopping criteria

$$\|\Gamma^{-1/2} (y - \bar{w}_n)\|_Y \approx \delta$$

An iterative regularizing ensemble Kalman method

Let $\rho < 1$ and $\tau > 1/\rho$. Generate an initial ensemble $u_0^{(j)} \sim \mu_0$

A regularizing Kalman method

(1) **Prediction Step:** Evaluate $w_m^{(j,f)} = \mathcal{G}(u_m^{(j)})$ define \bar{w}_m^f

(2) Stopping criteria. If

$$\|\Gamma^{-1/2}(y - \bar{w}_m^f)\| \leq \tau\eta$$

Stop. Otherwise: define C_m^{uw} , \bar{u}_m , C_m^{ww} and

(3) **Analysis step:** Compute the updated ensembles

$$u_{m+1}^{(j)} = u_m^{(j)} + C_m^{uw}(C_m^{ww} + \alpha_m\Gamma)^{-1}(y^{(j)} - w_m^{(j,f)})$$

for α_m such that

$$\alpha_m \|\Gamma^{1/2}(C_m^{ww} + \alpha_m\Gamma)^{-1}(y^\eta - \bar{w}_m^f)\| \leq \rho \|\Gamma^{-1/2}(y^\eta - \bar{w}_m^f)\|$$

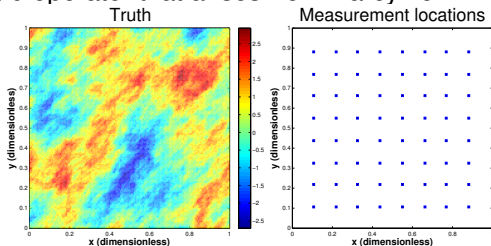
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Synthetic experiment with Darcy flow model

Initial ensemble generated from a prior $\mathbb{P}(u) = N(\bar{u}, C)$.
 $\mathcal{G}(u)$ be the forward operator that arises from Darcy flow.



Consider a truth $u^\dagger \sim \mathbb{P}(u)$ from which synthetic data are generated by $y = \mathcal{G}(u^\dagger) + \xi$ ($\xi \sim N(0, \Gamma)$ (prescribed Γ covariance of the Gaussian noise)).

For the numerical investigation with respect to the approximation properties of the Bayesian posterior see



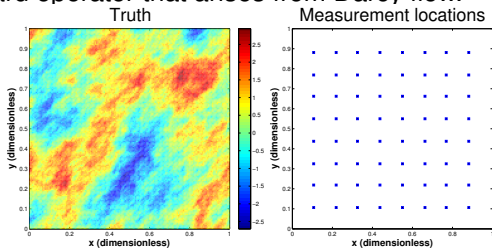
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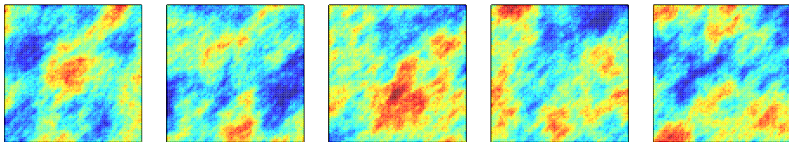
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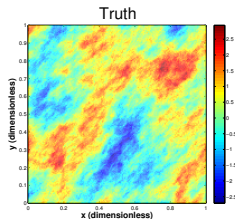


Some elements from the initial ensemble

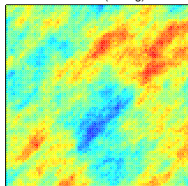


Results from the standard ES choice $\alpha = 1$.

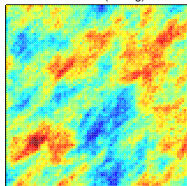
Reconstructing the truth with the mean of an ensemble of $N_e = 75$ (with small noise)



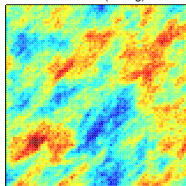
Ensemble mean (no reg). Iter:1



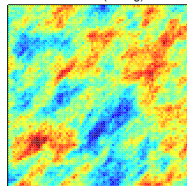
Ensemble mean (no reg). Iter:2



Ensemble mean (no reg). Iter:3



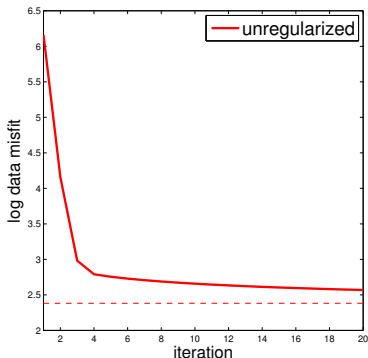
Ensemble mean (no reg). Iter:12



$$\bar{u}_n \equiv \frac{1}{N_e} \sum_{j=1}^{N_e} u_n^{(j)}$$

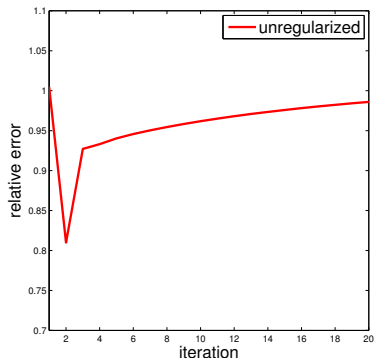
$$\|\Gamma^{-1/2}(y - \mathcal{G}(\bar{u}_n))\|_{\rho^2}$$

Data misfit



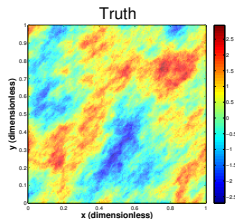
$$\|\bar{u}_n - u^\dagger\|_{L^2(D)}$$

Error w.r.t truth

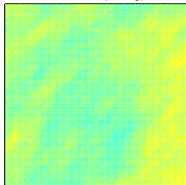


Results with the regularizing ensemble Kalman method

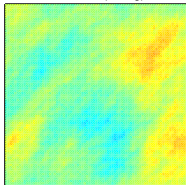
Reconstructing the truth with the mean of an ensemble of $N_e = 75$ (with small noise)



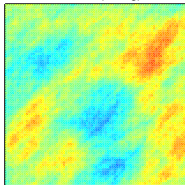
Ensemble mean (no reg). Iter:5



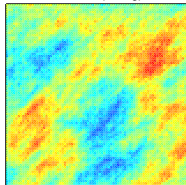
Ensemble mean (no reg). Iter:9



Ensemble mean (no reg). Iter:13



Ensemble mean (no reg). Iter:17

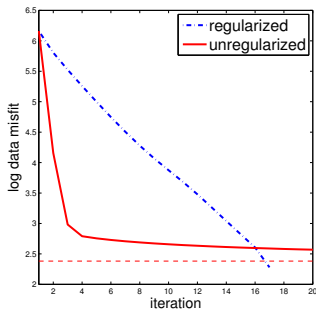


$$\bar{u}_n \equiv \frac{1}{N} \sum_{j=1}^N u_n^{(j)}$$

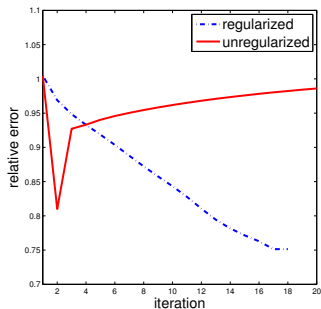
$$\|\Gamma^{-1/2}(y - \mathcal{G}(\bar{u}_n))\|_{\rho^2}$$

$$\|\bar{u}_n - u^\dagger\|_{L^2(D)}$$

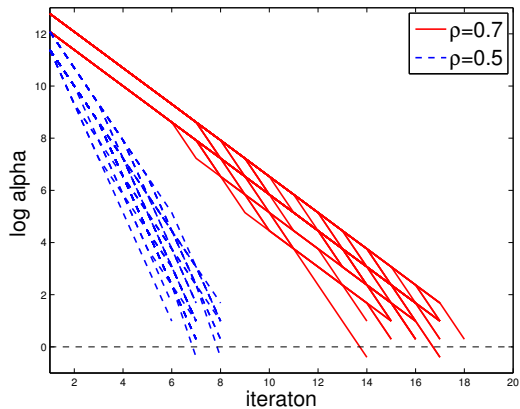
Data misfit



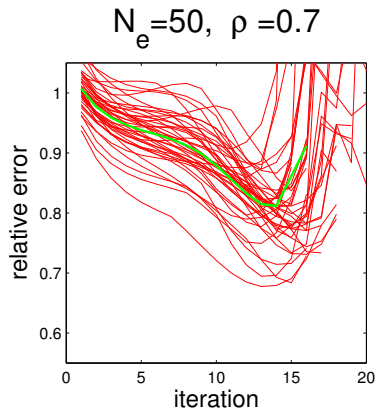
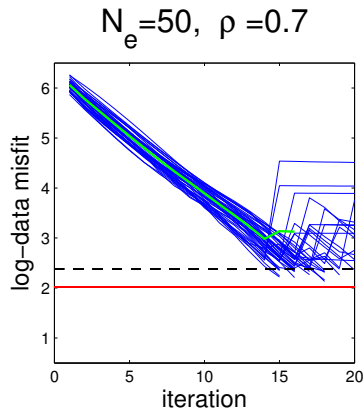
Error w.r.t truth



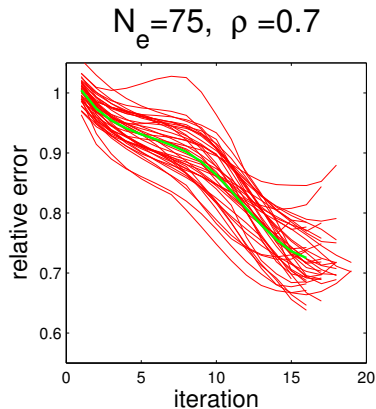
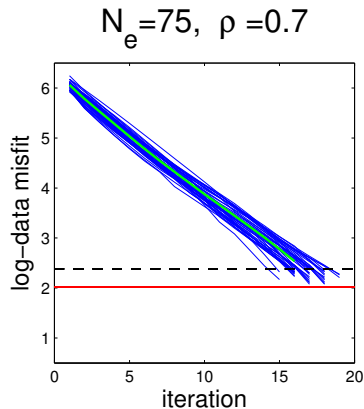
Plot of $\log \alpha$



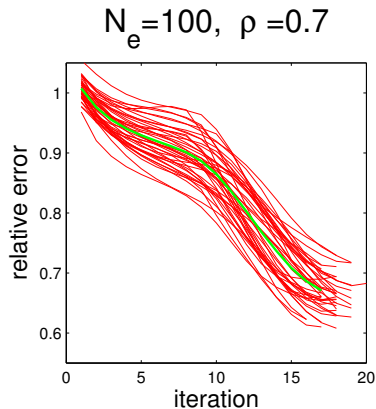
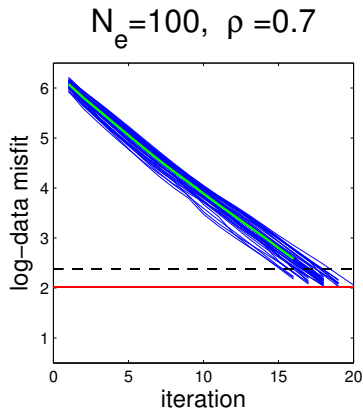
Regularizing properties as a function of the ensemble size



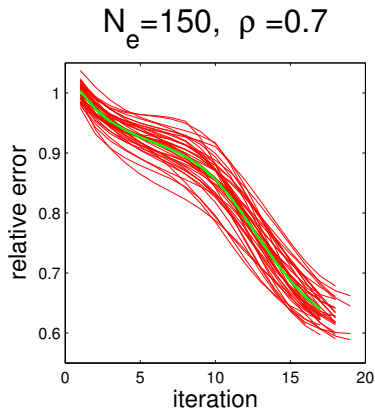
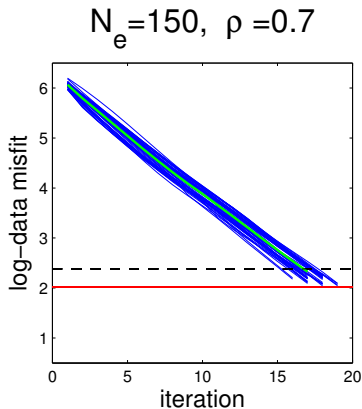
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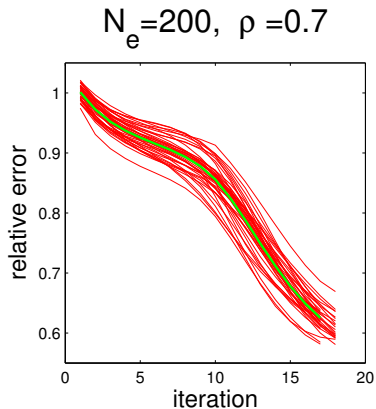
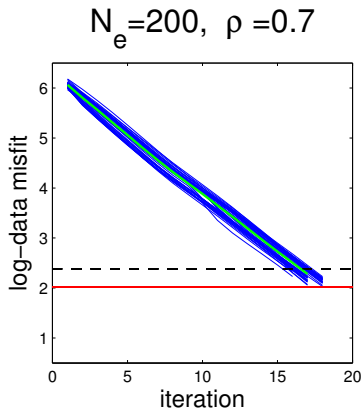
Regularizing properties as a function of the ensemble size



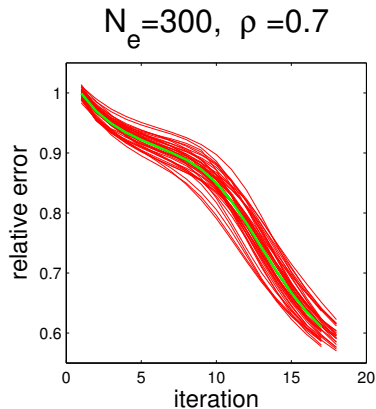
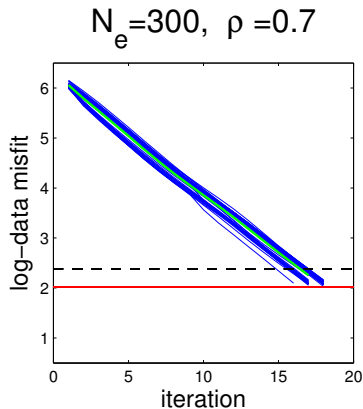
Regularizing properties as a function of the ensemble size



Regularizing properties as a function of the ensemble size



Regularizing properties as a function of the ensemble size

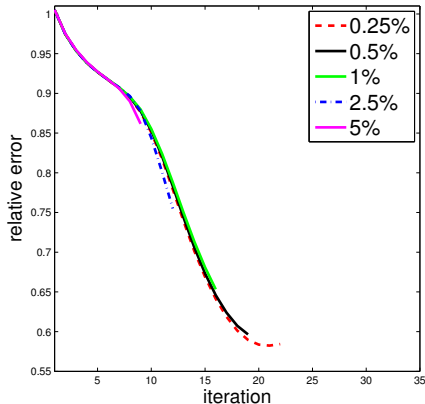
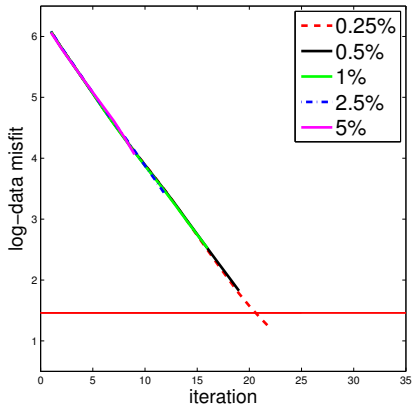


Convergence as the noise level decreases

$$\bar{u}_n \equiv \frac{1}{N_e} \sum_{j=1}^{N_e} u_n^{(j)}$$

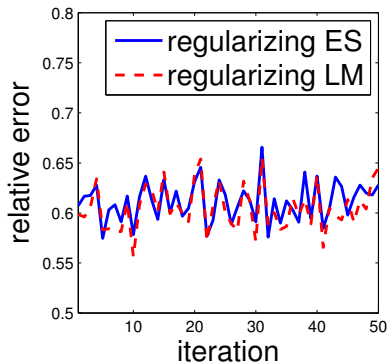
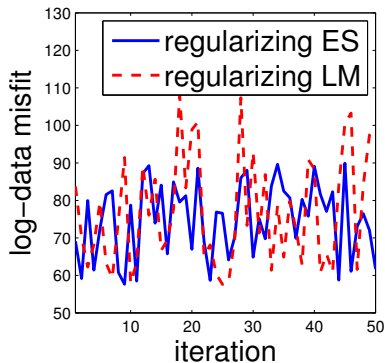
$$\|\Gamma^{-1/2}(y - \mathcal{G}(\bar{u}_n))\|_{\rho^2}$$

$$\|\bar{u}_n - u^\dagger\|_{L^2(D)}$$



The proposed ES as an approximate regularizing LM scheme

Comparing ES with the regularizing LM scheme (on the same subspace)



1 Introduction

2 Numerical Investigation of the Scheme

3 Applications

In Collaboration with Michael Tretyakov (UoN, maths) and Minh Park (UoN, maths), Mikhail Matveev (UoN, engineering)

Forward map: Resin Transfer Molding

$$\begin{aligned} -\nabla \cdot e^u \nabla p &= f && \text{in } D(t) \\ p &= p_{in} && \text{on } \Gamma_{in} \\ p &= p_f && \text{on } \Gamma_s(t) \\ -e^u \nabla p \cdot n &= 0 && \text{on } \Gamma_N \end{aligned}$$

Moving boundary

$$\frac{d\Gamma_s(t)}{dt} = -e^u \nabla p$$

$$u = \log(\kappa(x)) \in X \equiv L^\infty(D) \longrightarrow G(u) = \{p(x_i)\}_{i=1}^N \in Y \equiv \mathbb{R}^M$$

The Inverse Problem

Find $u \in X$ given

$$y = G(u) + \eta \quad \eta \sim N(0, \Gamma)$$

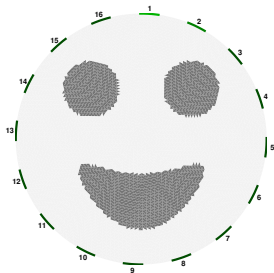
Complete Electrode Model: Forward and Inverse Problem

Given κ , $\{z_m\}_{m=1}^{n_e}$ and $I = \{I_m\}_{m=1}^{n_e}$ compute v and $V = \{V_m\}_{m=1}^{n_e}$

$$\begin{aligned}\nabla \cdot \kappa \nabla v &= 0 && \text{in } D, \\ v + z_m \kappa \nabla v \cdot \nu &= V_m && \text{on } e_m, \quad m = 1, \dots, n_e, \\ \nabla v \cdot \nu &= 0 && \text{on } \partial D \setminus \cup_{m=1}^{n_e} e_m, \\ \int_{e_m} \kappa \nabla v \cdot \nu \, ds &= I_m && m = 1, \dots, n_e,\end{aligned}$$

Inverse Problem:

Given $I^{(1)}, \dots, I^{(N)}$ and the observations of voltages $V^{(1)}, \dots, V^{(N)}$ find κ and z_m



Permeability κ defined through **level set function** u :

$$\kappa(x) = \kappa_1 \chi_{\{u < 0\}}(x) + \kappa_2 \chi_{\{u \geq 0\}}(x).$$

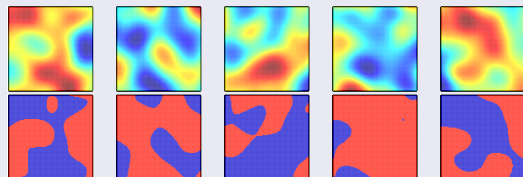
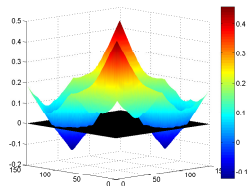
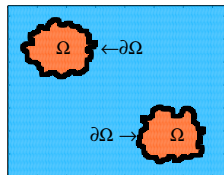
$u \mapsto \kappa$ is **discontinuous**

Forward Map and initial ensemble

$$u \longrightarrow \kappa \longrightarrow \mathcal{G}(u) = \{p(x_i)\}_{i=1}^N \in \mathbb{R}^M$$

Gaussian prior $\mu_0 = N(0, C_0)$ on the level-set function.

Covariance C_0 reflects the **regularity of the shape**.









M. A. Iglesias, Y. Lu and A. M. Stuart

A level-set approach to Bayesian geometric inverse problems.

Submitted, 2015. <http://arxiv.org/abs/1504.00313>

- Iterative regularization provides strategies for regularizing Kalman based methods.
- Regularization has strong effect in the robustness and accuracy of ensemble methods for solving both classical and Bayesian inverse problems.
- The stabilization of the proposed method is suitable for solving level-set based geometric inverse problems.
- Further investigations are required to establish the mathematical properties of these approximations.

-  M.A.Iglesias, K.Lin and A.M.Stuart
Well-posed Bayesian geometric inverse problems arising in subsurface flow.
Inverse Problems, 30 (2014) 114001
-  M. A.Iglesias
A regularizing ensemble Kalman method for PDE-constrained inverse problems.
to appear in Inverse Problems, 2015. <http://arxiv.org/abs/1505.03876>
-  M. A. Iglesias, Y. Lu and A. M . Stuart
A level-set approach to Bayesian geometric inverse problems.
Submitted, 2015. <http://arxiv.org/abs/1504.00313>
-  M. A.Iglesias
Iterative regularization for ensemble data assimilation in reservoir modeling
Computational Geosciences, (2015) 19:177-212
-  M. Iglesias, K. Law and A.M. Stuart,
Ensemble Kalman methods for inverse problems.
Inverse Problems. 29 (2013) 045001 <http://arxiv.org/abs/1209.2736>
-  S.L.Cotter, G.O.Roberts, A.M. Stuart, and D. White.
MCMC methods for functions: modifying old algorithms to make them faster.
Statistical Science, **28**(2013) 424-446). [arXiv:1202.0709](https://arxiv.org/abs/1202.0709).